

1. 补充证明.

1. 对任意 $n \geq 1$ 序列 X_n 证明:

$$\frac{S_n}{n} \xrightarrow{P} 0 \Rightarrow \frac{X_n}{n} \xrightarrow{P} 0.$$

$$X_n \xrightarrow{a.s.} 0 \Rightarrow \frac{S_n}{n} \xrightarrow{a.s.} 0$$

$$X_n \xrightarrow{L^p} 0 \Rightarrow \frac{S_n}{n} \xrightarrow{L^p} 0 \quad p \geq 1$$

举反例: $X_n \xrightarrow{P} 0 \not\Rightarrow \frac{S_n}{n} \xrightarrow{P} 0$.

证明:

由 Stolz 定理可得.

$$(1) \quad \frac{X_n}{n} = \frac{S_n - S_{n-1}}{n}$$

$$P\left(\left|\frac{X_n}{n}\right| > \varepsilon\right) = P\left(\left|\frac{S_n - S_{n-1}}{n}\right| > \varepsilon\right)$$

$$\leq P\left(\left|\frac{S_n}{n}\right| > \frac{\varepsilon}{2}\right) + P\left(\left|\frac{S_{n-1}}{n-1}\right| > \frac{\varepsilon}{2} \frac{n}{n-1}\right)$$

$\rightarrow 0$ as $n \rightarrow \infty$.

$$(3) \quad E\left|\frac{S_n}{n}\right|^p \leq \frac{\sum_{k=1}^n E|X_k|^p}{n} \xrightarrow{\text{Stolz}} \frac{E|X_n|^p}{1} \rightarrow 0. \quad \square$$

$$(4) \text{ (反例)} \quad X_n = \begin{cases} 2^n & \text{with prob. } \frac{1}{n} \\ 0 & \text{with prob. } 1 - \frac{1}{n} \end{cases}$$

$$nP(|X_n| > n) = 1 \not\rightarrow 0.$$

2. 若 $E|X| < \infty$ 证明: $\forall \varepsilon > 0, \sum_{n=1}^{\infty} 2^n P(|X| \geq 2^n \varepsilon) < \infty$

证: $\forall \varepsilon > 0$

$$\sum_{n=1}^{\infty} 2^n P(|X| \geq 2^n \varepsilon) = \sum_{k=1}^{\infty} 2^k P\left(\sum_{n=k}^{\infty} P(2^k \varepsilon \leq |X| < 2^{k+1} \varepsilon)\right)$$

$$= \sum_{k=1}^{\infty} 2^k \sum_{n=k}^{\infty} E 1_{\{2^k \varepsilon \leq |X| < 2^{k+1} \varepsilon\}}$$

$$\leq \sum_{k=1}^{\infty} 2^k \sum_{n=k}^{\infty} E\left[\frac{|X|}{2^k \varepsilon} 1_{\{2^k \varepsilon \leq |X| < 2^{k+1} \varepsilon\}}\right]$$

$$= \sum_{k=1}^{\infty} \sum_{n=k}^{\infty} E\left[\frac{|X|}{2^{k+n} \varepsilon} 1_{\{2^k \varepsilon \leq |X| < 2^{k+n} \varepsilon\}}\right]$$

$$= \sum_{k=1}^{\infty} \sum_{n=k}^{\infty} E\left[\frac{|X|}{2^{k+n} \varepsilon} 1_{\{2^k \varepsilon \leq |X| < 2^{k+n} \varepsilon\}}\right]$$

$$\leq \sum_{k=1}^{\infty} 2^k E\left[\frac{|X|}{\varepsilon} 1_{\{2^k \varepsilon \leq |X| < 2^{k+1} \varepsilon\}}\right]$$

$$= 2 \frac{E|X|}{\varepsilon} < \infty. \quad \square$$

3. X, X' iid. $\forall p > 0, E|X|^p < \infty \Leftrightarrow E|X - X'|^p < \infty$

证: \Rightarrow :

$$E|X - X'|^p \leq 2^p E|X|^p < \infty.$$

由对称性可得

\Leftarrow : $\forall x > 0$.

$$P(|X - X'| > x) \geq \frac{1}{2} P(|X - mX| \geq x).$$

$$\Rightarrow \int_0^{\infty} p \lambda^{p-1} P(|X - X'| \geq x) dx$$

$$\geq \frac{1}{2} \int_0^{\infty} p \lambda^{p-1} P(|X - mX| \geq x) dx$$

$$\Rightarrow \infty > E|X - X'|^p$$

$$\geq \frac{1}{2} E|X - mX|^p$$

$$\Rightarrow \frac{1}{2^p} E|X - mX|^p \leq \|X - X'\|_p < \infty.$$

$$X - mX \in L^p \Rightarrow Y \in L^p$$

4. 举例说明:

$$X_n \text{ iid. } \frac{S_n}{n} \xrightarrow{P} 0 \not\Rightarrow \frac{S_n}{n} \xrightarrow{a.s.} 0$$

证: X_n iid.

$$P(X_n = \pm n) = \frac{c}{n^2 \log n}, \quad n \geq 3.$$

$$c = \frac{1}{2 \sum_{n=3}^{\infty} \frac{1}{n^2 \log n}}$$

$$\frac{1}{2} \text{ 证 } X_n \xrightarrow{P} 0.$$

$$n E[X_n 1_{\{|X_n| > n\}}] = n \sum_{k > n} \frac{ck}{k^2 \log k}$$

$$\leq n \cdot \frac{1}{n \log n} \rightarrow 0.$$

$$\frac{1}{n^2} E[X_n^2 1_{\{|X_n| \leq n\}}] = \frac{1}{n} \sum_{k=3}^n \frac{ck^2}{k^2 \log k}$$

$$\leq \frac{1}{\log n} \rightarrow 0.$$

\therefore 由 Feller's Thm

$$\Rightarrow \frac{X_n}{n} \xrightarrow{P} 0$$

$$\Rightarrow \frac{S_n}{n} \xrightarrow{P} 0.$$

下证 $\not\Rightarrow$ a.s. $\rightarrow 0$

$$P(|X_n| > n) = O\left(\frac{1}{n \log n}\right) \Rightarrow \sum_{n=1}^{\infty} P(|X_n| > n) = \infty$$

由 Borel-Cantelli 引理 $P(|X_n| > n \text{ i.o.}) = 1$

$$|X_n| = |S_n - S_{n-1}|$$

$$\rightarrow P(|S_n| > \frac{n}{2} \text{ i.o.}) = 1 \Rightarrow \frac{S_n}{n} \not\xrightarrow{a.s.} 0 \quad \square$$

3.1) $k > p$; $\forall t > 0$

$$\lim_{n \rightarrow \infty} e^{-nt} \sum_{k=0}^{\infty} \frac{(nt)^k}{k!} = 1 \quad T > 0$$

14. F 是 \mathbb{R}^+ 上的 cdf. $\phi(t) = \int_0^{\infty} e^{-t\lambda} dF(\lambda)$

证: $\forall t > 0$ $\sum_{k=0}^{\infty} \frac{(nt)^k}{k!} \phi^{(k)}(t) \rightarrow F(1)$

证: $X_n \sim \text{Poi}(1)$, $E X_n = 1$, $S_n = X_{n+1} + \dots + X_n \sim \text{Poi}(n)$

由弱大数定律 $\frac{S_n - n}{n} \xrightarrow{P} 0 \Rightarrow \frac{S_n}{n} \xrightarrow{P} 1$

$T < t$ 时, $\exists \varepsilon > 0$, $\frac{T}{t} < 1 - \varepsilon$

$$e^{-nt} \sum_{k=0}^{\infty} \frac{(nt)^k}{k!} = P^*(t S_n \leq T n) \leq P\left(\frac{S_n}{n} \leq 1 - \varepsilon\right) \rightarrow 0 \text{ as } n \rightarrow \infty$$

$T > t$ 时, $\exists \varepsilon > 0$, $\frac{T}{t} < \varepsilon + 1$

$$\therefore P\left(\frac{S_n}{n} \leq 1 - \varepsilon\right) < P\left(\frac{S_n}{n} \leq \frac{T}{t}\right) \leq P\left(\frac{S_n}{n} \leq \varepsilon + 1\right)$$

$$n \rightarrow \infty \text{ 有 } P\left(\frac{S_n}{n} \leq \varepsilon + 1\right) \rightarrow 1$$

$$(2) T_n = \sum_{k=0}^{\infty} \frac{(nt)^k}{k!} \phi^{(k)}(t) = \int_0^{\infty} \sum_{k=0}^{\infty} e^{-nt} \frac{(nt)^k}{k!} dF(t) = \int_0^T \sum_{k=0}^{\infty} e^{-nt} \frac{(nt)^k}{k!} dF(t) + \int_T^{\infty} \sum_{k=0}^{\infty} e^{-nt} \frac{(nt)^k}{k!} dF(t)$$

$$T_n - F(T) = \int_0^T \left(\sum_{k=0}^{\infty} e^{-nt} \frac{(nt)^k}{k!} - 1 \right) dF(t)$$

$$+ \int_T^{\infty} \sum_{k=0}^{\infty} e^{-nt} \frac{(nt)^k}{k!} dF(t) \Rightarrow 1 - 1 + 0 = 0 \text{ as } n \rightarrow \infty$$

4. X, X_1, Y_2, \dots iid. 非负. 若极限

$$\lim_{s \rightarrow \infty} \frac{E[X \mathbb{1}_{\{X \leq s\}}]}{s P(X > s)} = \infty$$

$\{a_n\}$ s.t. $\frac{a_n}{a_{n+1}} \rightarrow 1$

证明: $\forall \varepsilon > 0$ $\exists M(s) = E[X \mathbb{1}_{\{X \leq s\}}]$

$$V(s) = \frac{M(s)}{s P(X > s)} \rightarrow \infty$$

$$\forall \frac{M(s)}{s} = \int_0^s \frac{1}{s} dF(x) \frac{P(X \leq x)}{s}$$

$$\therefore \forall n \exists S_n, S_n \rightarrow \infty, \frac{M(S_n)}{S_n} < \frac{1}{n}$$

$$n P(X_n > S_n) = \frac{S_n (1 - F(S_n))}{\mu(S_n)} = \frac{1}{F(S_n)} \rightarrow 0$$

$$\frac{1}{S_n} n E[X_n \mathbb{1}_{\{X_n \leq S_n\}}]$$

$$= \frac{\int_0^{S_n} 2x P(X > x) dx}{\int_0^{S_n} \mu(x) dx} \rightarrow 0$$

$$\therefore \frac{S_n}{S_n} \xrightarrow{P} 1$$

5. $\{X, X_n\}$ iid. $\forall x > 0, P(|X| > x) < \infty$

证: $\forall n \in \mathbb{N}, \lim_{x \rightarrow \infty} \frac{P(\max_{1 \leq k \leq n} |X_k| > x)}{n P(|X| > x)} = 1$

$$\begin{aligned} \underline{\text{证}}: P(\max_{1 \leq k \leq n} |X_k| > x) &= 1 - P(|X_1| \leq x, \dots, |X_n| \leq x) \\ &= 1 - P(|X| \leq x)^n \\ &= P(|X| > x) (1 + P(|X| \leq x) + \dots + P^{n-1}(|X| \leq x)) \\ &\Rightarrow \frac{1}{x} \rightarrow \frac{n}{n} = 1 \end{aligned}$$

6. $X_n \xrightarrow{P} 0$. 证: $\liminf_{n \rightarrow \infty} |X_n| = 0$ a.s.

证: $X_n \xrightarrow{P} 0, \exists \text{ 对 } \forall \varepsilon > 0, X_{n_k} \xrightarrow{a.s.} 0$

$\exists \varepsilon > 0, \forall n, \exists k, n_k \leq n < n_{k+1}$

$$0 \leq \liminf_{n \rightarrow \infty} |X_n|$$

$$\square \quad = \liminf_{n \rightarrow \infty} |X_n| \leq \liminf_{j \rightarrow \infty} |X_{n_j}|$$

$$= \liminf_{j \rightarrow \infty} |X_{n_j}| = 0 \text{ a.s.}$$

6. 考虑 \mathbb{Z}^d 上独立随机游动. $\{X_n\}$ i.i.d. 且 X_1 从

$\{(a_1, \dots, a_d) : a_i = \pm 1\}$ 上均匀分布. $S_n = \sum_{i=1}^n X_i$

证明: $d \geq 3$ 时, $P(S_n = 0, i.o.) = 0$.

证: $S_n = 0 \iff$ 每个分量都为 0 $\iff n = 2m, m \in \mathbb{Z}_+$.

$$\sum_{m=1}^{\infty} P(S_{2m} = 0)$$

$$= \sum_{m=1}^{\infty} \binom{2m}{m} 2 \cdot \frac{1}{2^{2m}}^d$$

$$= \sum_{m=1}^{\infty} \frac{(2m)!}{m! n!} \frac{1}{2^{2m-1}}$$

\Rightarrow 由 Stirling 公式 $n! \sim \sqrt{2\pi n} \left(\frac{n}{e}\right)^n$

\therefore 证 (1) 收敛于

$$\sum_{m=1}^{\infty} \left(\frac{\sqrt{4\pi m} \cdot e^{-2m}}{2^{2m-1}} \cdot \frac{2^{2m-1}}{\sqrt{2\pi m}} \right)^d$$

$$\sim \sum_{m=1}^{\infty} \frac{1}{m^{d/2}} < \infty \quad (d \geq 3)$$

$$\therefore \sum_{m=1}^{\infty} P(S_{2m} = 0) < \infty \Rightarrow P(S_{2m} = 0 \text{ i.o.}) = 0$$

$$\text{又 } P(S_{2m-1} = 0) = 0$$

$$\therefore P(S_n = 0 \text{ i.o.}) = 0.$$

7. X, X_1, \dots, X_n i.i.d. 非负.

$$\text{证明: (1) } \limsup_{n \rightarrow \infty} \frac{X_n}{n} = 0 \text{ 或 } EX < \infty$$
$$\left\{ \begin{array}{l} \text{或} \\ \omega \end{array} \right. \quad EX = \infty$$

(2) 证明: 当 $EX < \infty$ 时, 对任意 $0 < c < 1$,

都有 $\sum_n e^{X_n c^n} < \infty$ a.s.

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证: (1) $EX < \infty$ 时

$$\forall A > 0 \quad EX < \infty \Rightarrow \sum_{n=1}^{\infty} P(X_n > A^n) < \infty$$

$$\Rightarrow \sum_{n=1}^{\infty} P(X_n > A^n \text{ i.o.}) < \infty$$

$$\Rightarrow \limsup_{n \rightarrow \infty} \frac{X_n}{n} = 0$$

$EX < \infty$ 时

$$\forall A > 0 \quad E\left[\frac{X}{A}\right] < \infty \Rightarrow \sum_{n=1}^{\infty} P(X_n > A^n \text{ i.o.}) < \infty$$

$$\Rightarrow \limsup_{n \rightarrow \infty} \frac{X_n}{n} = 0$$

$$(2) \quad e^{X_n c^n} = (e^{\frac{X_n}{c}})^n$$

$$\text{若 } EX < \infty \text{ 则 } \limsup_{n \rightarrow \infty} \frac{X_n}{n} = 0 \in \liminf_{n \rightarrow \infty} \frac{X_n}{n} \Rightarrow \frac{X_n}{n} \rightarrow 0 \text{ a.s.}$$

$$\therefore \forall \varepsilon \in (0, \log \frac{1}{2c}). \exists N \text{ s.t. } \forall n > N \quad \frac{X_n}{n} < \varepsilon$$

$$\Rightarrow \sum_{n=1}^{\infty} (e^{\frac{X_n}{c}} c)^n \leq \sum_{n=1}^N (e^{\frac{X_n}{c}} c)^n + \sum_{n=N+1}^{\infty} (e^{\frac{X_n}{c}} c)^n$$

$$\leq N + \frac{2}{1-c} < \infty$$

$$EX < \infty \text{ 时, } \limsup_{n \rightarrow \infty} \frac{X_n}{n} = 0$$

$$\exists R. P(|Z| = 1) = 1 \text{ s.t. } \forall \varepsilon \in \mathbb{Q}. \exists \eta_{\varepsilon}(\omega)$$

$$\frac{X_{n_k}(\omega)}{n_k(\omega)} > A$$

$$\Rightarrow \sum_{k=1}^{\infty} e^{X_{n_k} c^{n_k}} > \sum_{k=1}^{\infty} (e^A c)^{n_k} = \infty \text{ a.s.}$$

8. X_n 独立 $\sim \text{Poi}(\lambda_n)$.

$$\text{证明: 若 } \sum_n \lambda_n = \infty \text{ 则 } \frac{S_n}{E S_n} \rightarrow 1 \text{ a.s.}$$

证: 不妨设 $\lambda_n < 1$

由 Chebyshev 不等式

$$P(|S_n - E S_n| > \delta E S_n)$$

$$\leq \frac{\text{Var}(S_n - E S_n)}{\delta^2 E^2 S_n} = \frac{1}{\delta^2 E S_n} \rightarrow 0 \text{ as } n \rightarrow \infty$$

$$\therefore \frac{S_n}{E S_n} \xrightarrow{P} 1$$

$$\text{取 } \omega \in \eta_{\varepsilon}(\omega) \text{ 的 } \{n: E S_n \geq k^2\} \quad EX_n \leq 1$$

$$\text{则 } k^2 \leq E T_k \leq k^2 + 1$$

$$\therefore P(|T_k - E T_k| > \sqrt{E T_k}) \leq \frac{1}{\delta^2 k^2}$$

求和 $\sum_{k=1}^{\infty} \frac{1}{k^2} < \infty$, 由 Borel-Cantelli 引理

$$\frac{T_k}{E T_k} \xrightarrow{a.s.} 1$$

$$n_k \leq n < n_{k+1}$$

$$\frac{E T_k}{E T_{k+1}} = \frac{T_k}{E T_{k+1}} \leq \frac{S_n}{E S_n} \leq \frac{T_{k+1}}{E T_k} = \frac{T_{k+1}}{E T_{k+1}} \cdot \frac{E T_{k+1}}{E T_k}$$

$$k \rightarrow \infty$$

$$\frac{S_n}{E S_n} \rightarrow 1 \text{ a.s.}$$

□

9. $f, g: \mathbb{R} \rightarrow \mathbb{R}$ 单增 $\sum_{i=1}^n x_i$ 且 $\forall n, X$

$$E f(X) g(X) \geq E f(X) E g(X)$$

$$\Downarrow$$

$$\text{Cov}(f, g) \geq 0$$

证: 由 $f \uparrow$

$$\therefore (f(x) - f(x'))(g(x) - g(x')) \geq 0$$

取 x, x' 即有 (在末标号间)

$$E \frac{1}{n} \sum_{i=1}^n (f(x_i) - f(x_{i-1}))(g(x_i) - g(x_{i-1})) \geq 0$$

□

10. 独立投 n 枚硬币, 正向出现的概率为 p .

$$A_k = \{ \text{第 } 2^k \rightarrow 2^{k+1} \text{ 次中间存在连续 } k \text{ 次正向} \}$$

$$\text{证: } P(A_k \text{ i.o.}) = \begin{cases} 0 & p < \frac{1}{2} \\ 1 & p > \frac{1}{2} \end{cases}$$

证: (1) $p < \frac{1}{2}$

$$E_i = \{ \text{从 } 2^k + i - 1 \text{ 次开始连续 } k \text{ 次正向} \}$$

$$A_k = \bigcup_{i=1}^{2^k - k} E_i$$

$$P(A_k) \leq \sum_{i=1}^{2^k - k} P(E_i) = p^k (2^k - k)$$

$$\sum_{k=1}^{\infty} (2^k - k) p^k \leq \frac{1}{1-2p} < \infty$$

(2) $p > \frac{1}{2}$

$$P(A_k) = P(\underbrace{E_1 \cup E_{k+1} \cup \dots \cup E_{k+1}}_{\text{这些互斥}}) \quad l = \lfloor \frac{2^{k+1} - k}{k} \rfloor$$

$$= 1 - P(E_1^c \cap \dots \cap E_{k+1}^c)$$

$$= 1 - P(E_1^c)^{l+1} = (-(1-p^k))^{l+1} \geq 1 - (1 - (\frac{1}{2})^k)^{l+1}$$

$$\sum_{k=1}^{\infty} \geq 1 - \exp(-\frac{1}{2}) \sim \frac{1}{2}$$

求和 $= \infty$

11. 考虑 Z^2 每步以概率 p 开

$1-p$ 关 独立

□

把关这步, 记: 存在一个连续的开过程概率为 0 或 1.

pf. 因这是尾事件

□

12. $\{X_n\}$ iid. 分布函数 $F(x)$ 连续. $Y_n = F^{-1}(X_n)$

证明: (1) $Y_n \rightarrow 1$ a.s.

$$(2) \sum_{n=1}^{\infty} (1 - Y_n) \text{ a.s. 收敛}$$

证: (1) $Y_n = F^{-1}(X_n)$, $0 \leq Y_n \leq 1$.

$$\forall \epsilon > 0, P(|Y_n - 1| \geq \epsilon)$$

$$= P(Y_n \leq 1 - \epsilon)$$

$$= P(F(X_n) \leq (1 - \epsilon)^n)$$

$$= P(X_n \leq F^{-1}((1 - \epsilon)^n)) = (1 - \epsilon)^n$$

$$\therefore \sum_{n=1}^{\infty} P(|Y_n - 1| \geq \epsilon) = \sum_{n=1}^{\infty} (1 - \epsilon)^n < \infty$$

\therefore 由 Borel-Cantelli 引理.

$$P(|Y_n - 1| \geq \epsilon \text{ i.o.}) = 0, \forall \epsilon > 0$$

$$\therefore Y_n \xrightarrow{\text{a.s.}} 1$$

(2) $1 - Y_n = 1 - EY_n + EY_n - Y_n$

$$EY_n = E[F^{-1}(X_n)]$$

$$= \int F^{-1}(x) dF(x)$$

$$= \frac{n}{n+1}$$

$$\therefore \sum_{n=1}^{\infty} (1 - EY_n) \geq \sum_{n=2}^{\infty} \frac{1}{n} = \infty$$

~~EY_n~~ 对 $EY_n - Y_n$

$$E[EY_n - Y_n] = 0$$

$$\text{又 } EY_n^2 = \int_{-\infty}^{\infty} F^{-2}(x) dF(x)$$

$$= \frac{n}{n+2} F^{-\frac{n}{n+2}}(x) \Big|_{-\infty}^{+\infty} = \frac{n}{n+2}$$

$$\therefore \sum_{n=1}^{\infty} \text{Var}(EY_n - Y_n)$$

$$= EY_n^2 - E^2 Y_n$$

$$= \frac{n}{n+2} - \frac{n^2}{(n+1)^2}$$

$$= \frac{n}{(n+2)(n+1)^2} < \frac{1}{n^2}$$

求和 $< \infty$, 由 Borel-Cantelli 引理.

$$\sum_{n=1}^{\infty} EY_n - Y_n \text{ 收敛}$$

\therefore 证得收敛

13. 设 $\{X_n\}$ 独立

(1) 若 $\sum_{n=1}^{\infty} X_n^2 < \infty$ a.s. 则 $\sum_{n=1}^{\infty} X_n$ 收敛 a.s. $\Leftrightarrow \sum_{n=1}^{\infty} E[X_n I_{\{|X_n| \leq 1\}}]$ 收敛

(2) 若 $\sum_{n=1}^{\infty} X_n$ a.s. 收敛 则 $\sum_{n=1}^{\infty} E[X_n^2] < \infty$ a.s. $\Leftrightarrow \sum_{n=1}^{\infty} E[X_n^2 I_{\{|X_n| \leq 1\}}] < \infty$

证: (1) \Rightarrow 若 $\sum_{n=1}^{\infty} X_n^2 < \infty$ a.s. 则 $\sum_{n=1}^{\infty} X_n$ a.s. 收敛. 由三收敛定理
 $\sum_{n=1}^{\infty} E[X_n I_{\{|X_n| \leq 1\}}]$ 收敛
 $\Leftrightarrow \sum_{n=1}^{\infty} E[X_n^2 I_{\{|X_n| \leq 1\}}] < \infty$
 $\sum_{n=1}^{\infty} X_n^2 < \infty$ a.s.
 由三收敛定理
 $\sum_{n=1}^{\infty} P(|X_n|^2 > 1) < \infty$ a.s.
 而 $|X_n|^2 > 1 \Leftrightarrow |X_n| > 1$
 $\therefore \sum_{n=1}^{\infty} P(|X_n| > 1) < \infty$
 再: $\sum_{n=1}^{\infty} E[|X_n|^2 I_{\{|X_n|^2 \leq 1\}}]$ 收敛
 \Downarrow
 $\sum_{n=1}^{\infty} E[|X_n|^2 I_{\{|X_n| \leq 1\}}]$ 收敛
 \Downarrow
 $\sum_{n=1}^{\infty} \text{Var}[X_n I_{\{|X_n| \leq 1\}}]$ 收敛

由三收敛定理即有结论
 (2) \Leftarrow $\sum_{n=1}^{\infty} X_n$ a.s. 收敛. 由三收敛定理
 $\sum_{n=1}^{\infty} P(|X_n| > 1) < \infty$
 $\sum_{n=1}^{\infty} E[X_n I_{\{|X_n| \leq 1\}}]$ 收敛
 $\sum_{n=1}^{\infty} \text{Var}(X_n I_{\{|X_n| \leq 1\}}) < \infty$

因 $\forall \omega \in \Omega, |X_n(\omega)| > 1 \Leftrightarrow |X_n(\omega)|^2 > 1$
 $\therefore \sum_{n=1}^{\infty} P(|X_n|^2 > 1) < \infty$
 $E[X_n^2 I_{\{|X_n| \leq 1\}}] = E[(X_n I_{\{|X_n| \leq 1\}})^2]$
 $= \text{Var}(X_n I_{\{|X_n| \leq 1\}}) + (E[X_n I_{\{|X_n| \leq 1\}}])^2$
 非和 $< \infty$
 又 $\text{Var}(X_n^2 I_{\{|X_n|^2 \leq 1\}}) \leq E[X_n^4 I_{\{|X_n| \leq 1\}}]$
 $\leq E[X_n^2 I_{\{|X_n| \leq 1\}}]$
 非和 $< \infty$

由Kolmogorov收敛定理知收敛
 $\Rightarrow \sum_{n=1}^{\infty} E[X_n^2 I_{\{|X_n| \leq 1\}}] < \infty \Rightarrow \sum_{n=1}^{\infty} E[X_n^2 I_{\{|X_n| \leq 1\}}] = E[\sum_{n=1}^{\infty} E[X_n^2 I_{\{|X_n| \leq 1\}}]] < \infty$

14. $\{X_n; n \geq 1\}$ 同分布. 对 $p > 0$.
 $E|X_1|^p < \infty$ 证: $\frac{X_n}{n^p} \rightarrow 0$ a.s.

证明: $\forall \epsilon > 0$
 有 $P(|\frac{X_n}{n^p}| \geq \epsilon)$
 $= P(|X_n| \geq \epsilon n^p)$
 $= P(|X_1|^p \geq \epsilon^p n)$
 非和 $< \infty$ 同 $E|X_1|^p < \infty$
 再由B-C引理即得. \square

15. $\{X, X_n; n \in \mathbb{Z}_+\}$ iid. $P(X=1) = \frac{1}{2}$
 $U_n = \sum_{k=1}^n \frac{X_k}{2^k}, n \geq 1$. 证明: $U_n \rightarrow U$ a.s. 其中 $U \sim U[-1, 1]$
 证明: $\{U_n\}$ 绝对收敛. $\forall \omega \in \Omega$
 $U_n \xrightarrow{\text{a.s.}} U$ 下证: $U \sim U[-1, 1]$

$\phi_U(t) = E[e^{itU}]$
 $\stackrel{\text{a.s.}}{=} \lim_{n \rightarrow \infty} E[e^{itU_n}]$
 $= \lim_{n \rightarrow \infty} E[e^{i \sum_{k=1}^n \frac{X_k}{2^k} t}]$
 $= \lim_{n \rightarrow \infty} \prod_{k=1}^n E[e^{i \frac{X_k}{2^k} t}]$
 $= \lim_{n \rightarrow \infty} \prod_{k=1}^n \frac{e^{i \frac{t}{2^k}} + e^{-i \frac{t}{2^k}}}{2}$
 $= \lim_{n \rightarrow \infty} \prod_{k=1}^n \cos \frac{t}{2^k}$
 $= \frac{\sin t}{t}$ \square

16. $\{X, X_n; n \geq 1\}$ iid. 证: $\frac{\log_n \sum_{k=1}^n X_k}{\log n} \xrightarrow{\text{a.s.}} 0 \Leftrightarrow E|X| < \infty$
 $E X = 0$
 证明: \Leftarrow 令 $Y_k = \frac{X_k}{\log k}, a_n = \frac{n}{\log n}$
 证明:

- ① $\sum_{n=1}^{\infty} P(|Y_n| > a_n) < \infty$
- ② $\sum_{n=1}^{\infty} E[Y_n^2 I_{\{|Y_n| \leq a_n\}}] < \infty$
- ③ $\sum_{n=1}^{\infty} \text{Var}[Y_n I_{\{|Y_n| \leq a_n\}}] < \infty$

$\Rightarrow \sum_{n=1}^{\infty} E[Y_n^2 I_{\{|Y_n| \leq a_n\}}] < \infty \Rightarrow \sum_{n=1}^{\infty} E[Y_n^2 I_{\{|Y_n| \leq a_n\}}] = E[\sum_{n=1}^{\infty} E[Y_n^2 I_{\{|Y_n| \leq a_n\}}]] < \infty$

(3) $P(|Y_n| > a_n)$

$= P\left(\left|\frac{X_n}{\log n}\right| > \frac{a_n}{\log n}\right)$

$= P(|X_n| > a_n)$

$\sum_{n=1}^{\infty} P(|Y_n| > a_n) = \sum_{n=1}^{\infty} P(|X_n| > a_n) < \infty$

(3) $\text{Var}\left[\frac{Y_n}{a_n} I_{\{|Y_n| \leq a_n\}}\right]$

$\leq E\left[\frac{Y_n^2}{a_n^2} I_{\{|Y_n| \leq a_n\}}\right]$

$= E\left[\frac{X_n^2}{n^2} I_{\{|X_n| \leq n\}}\right] = E\left[\frac{Y_n^2}{n^2} I_{\{|X_n| \leq n\}}\right]$

\therefore 求和 $\rightarrow E\left[\sum_{n=1}^{\infty} \frac{Y_n^2}{n^2}\right]$

$\leq E|X| < \infty$

剩下用(2). 即 $E\left[\frac{Y_n}{a_n} I_{\{|Y_n| \leq a_n\}}\right] \xrightarrow{a.s.} 0$

$Z_n = Y_n I_{\{|Y_n| \leq a_n\}}$

由(2) $\sum_{n=1}^{\infty} \text{Var} Z_n < \infty$ 又 $E[Z_n] = 0$

\therefore 由一级数定理 $\sum_{n=1}^{\infty} \frac{Z_n - EZ_n}{n} \xrightarrow{a.s.} 0$

由Kronecker引理 $\sum_{k=1}^n \frac{Z_k - EZ_k}{a_n} \xrightarrow{a.s.} 0$

$\therefore \sum_{k=1}^n \frac{Z_k}{a_n} \rightarrow 0 \text{ a.s.}$

$\sum_{n=1}^{\infty} P(Y_n + Z_n) = \sum_{n=1}^{\infty} P(|Y_n| > a_n) < \infty$

$\therefore \sum_{n=1}^{\infty} \frac{Y_n}{a_n} \xrightarrow{a.s.} 0$

$\Rightarrow \sum_{k=1}^n \frac{\log k}{n} \sum_{k=1}^n \left(\frac{X_k}{\log k}\right) \xrightarrow{a.s.} 0$

$\Rightarrow P\left(\left|\frac{Y_n}{n/\log n}\right| > \varepsilon \text{ i.o.}\right) = 0 \quad \forall \varepsilon > 0$

$\therefore \sum_{k=1}^{\infty} P\left(\left|\frac{Y_k}{n/\log n}\right| > \varepsilon\right) < \infty$

$\Leftrightarrow E|Y_n/\log n| = E|X_n|$

欲证 $EX=0$ 只需证 $E|X| < \infty$

则由 $E[X_k - EX_k] = 0$

$E|X_n - EX_n| < \infty \Rightarrow \frac{\log n}{n} \sum_{k=1}^n \frac{X_k - EX_k}{\log k} \xrightarrow{a.s.} 0$

$\therefore \frac{\log n}{n} \sum_{k=1}^n \frac{X_k}{\log k} \xrightarrow{a.s.} \frac{\log n}{n} \sum_{k=1}^n \frac{EX_k}{\log k}$

16. $\{X_n, n \in \mathbb{N}\}$ iid. $EX=0, E[|X| \log^+ |X|] < \infty$

其中 $\log^+ t = \ln\{t \vee e\}$

求证: $\frac{\log^+ n}{n} \sum_{k=1}^n X_k \xrightarrow{a.s.} 0$

证: $\varphi(x) = |x| \log^+ |x|, \varphi(x)$ 为 \mathbb{R} 上的凸函数

即 φ 满足 $\varphi(x) \geq 0$

$Y_n = X_n I_{\{|X_n| \leq a_n\}}, a_n = \frac{n}{\log^+ n}$

$\sum_{n=1}^{\infty} P(Y_n \neq X_n) = \sum_{n=1}^{\infty} P(|X_n| > a_n)$

$= \sum_{n=1}^{\infty} P(|X_n| > \frac{n}{\log^+ n})$

$= \sum_{n=1}^{\infty} P\left(\varphi(|X_n|) > \varphi\left(\frac{n}{\log^+ n}\right)\right)$

$= \sum_{n=1}^{\infty} P\left(|X| \log^+ |X| > n - n \frac{\log^+ \log^+ n}{\log^+ n}\right)$

$\leq \sum_{n=1}^{\infty} P(|X| \log^+ |X| > \frac{n}{2})$

$= \sum_{n=1}^{\infty} P(|X| \log^+ |X| > \frac{n}{2}) < \infty$

\therefore 只须证 $\frac{1}{a_n} \sum_{k=1}^n Y_k \xrightarrow{a.s.} 0$ 由Kronecker引理

为证:

(1) $\sum_{n=1}^{\infty} P(|Y_n| > a_n) = 0$

(2) $\sum_{k=1}^{\infty} \text{Var}\left(\frac{Y_k}{a_n}\right) \leq \sum_{k=1}^{\infty} \frac{E Y_k^2}{a_n^2} \leq \frac{E|X| \log^+ |X|}{a_n} < \infty$

(3) $\sum_{n=1}^{\infty} E\left[\frac{Y_n}{a_n}\right] = \sum_{n=1}^{\infty} \left[\frac{1}{a_n} E[X I_{\{|X| \leq a_n\}}]\right]$

$\stackrel{EX=0}{=} \sum_{n=1}^{\infty} \left[\frac{1}{a_n} E[X I_{\{|X| > a_n\}}]\right]$

$\leq \sum_{n=1}^{\infty} \frac{1}{a_n} E[|X| I_{\{|X| > a_n\}}]$

$\leq E\left[\sum_{n \geq |X|} \frac{|X|}{a_n}\right]$

$\leq E\left[\frac{|X|}{|X| \log^+ |X|}\right] = E[|X| \log^+ (|X| \log^+ |X|)]$

$\leq E[|X| \log^+ |X|] < \infty$

由三级数定理即得

2. 补充习题:

1. 证明 Lindeberg-Feller 定理.

t.f.a.e.
 (1) $\sum_{k=1}^{k_n} X_{n,k} \xrightarrow{d} N(0,1)$, 且 $\max_{1 \leq k \leq k_n} E[X_{n,k}^2] \rightarrow 0$.

(2) $\forall \varepsilon > 0, \sum_{k=1}^{k_n} E[X_{n,k}^2 1_{\{|X_{n,k}| \geq \varepsilon\}}] \rightarrow 0$.

注: 必须用引理.

$\{X_{n,k}\}$ 为独立序列, $Z \sim N(0,1)$.

且 $\sup_k P(|X_{n,k}| > \varepsilon) \rightarrow 0, \forall \varepsilon > 0, k_j$

$\sum_{k=1}^{k_n} X_{n,k} \xrightarrow{d} Z. (\Leftrightarrow)$

(1) $\sum_{k=1}^{k_n} P(|X_{n,k}| > \varepsilon) \rightarrow 0, \forall \varepsilon > 0$

(2) $\sum_{k=1}^{k_n} E[X_{n,k}^2 1_{\{|X_{n,k}| \leq 1\}}] \rightarrow b$

(3) $\sum_{k=1}^{k_n} \text{Var}[X_{n,k} 1_{\{|X_{n,k}| \leq 1\}}] \rightarrow c$.

2. X_1, X_2, \dots, X_n iid. X 对称.

$P(|X| > x) = \frac{1}{x^2}, x > 1$. 证明: $\frac{S_n}{\sqrt{n \log n}} \xrightarrow{d} N(0,1)$

证: 令 $X_{n,k} = \frac{X_k}{a_n} = \frac{X_k}{\sqrt{n \log n}}$.
 要证: $\sum_{k=1}^{k_n} X_{n,k} \xrightarrow{d} N(0,1)$.

无穷小条件:

$P(|X_{n,k}| > \varepsilon) = P(|X_k| > \varepsilon \sqrt{n \log n})$
 $= \frac{1}{\varepsilon^2 n \log n} \rightarrow 0$.

对 $\forall \varepsilon > 0$ 对 k -致成立

(1) $\forall \varepsilon > 0, \sum_{k=1}^{k_n} P(|X_{n,k}| > \varepsilon) \rightarrow 0$ ($k_n = n$).
 $= n \cdot P(|X_n| > \varepsilon \sqrt{n \log n}) = \frac{1}{\varepsilon^2 \log n} \rightarrow 0$.

(2) $\sum_{k=1}^n E[X_{n,k}^2 1_{\{|X_{n,k}| \leq 1\}}] = n E[X^2 1_{\{|X| \leq \sqrt{n \log n}\}}]$
 $\xrightarrow{\sqrt{n \log n}}$
 $= 0$. 因 X 对称.

(3) $\sum_{k=1}^n E[X_{n,k}^2 1_{\{|X_{n,k}| \leq 1\}}] = \frac{n}{n \log n} E[X^2 1_{\{|X| \leq \sqrt{n \log n}\}}]$
 $= \frac{1}{\log n} \int_1^{\sqrt{n \log n}} x^2 \frac{d}{dx} (1 - \frac{1}{x^2}) dx = \frac{\log n \log n}{\log n} \rightarrow 1$.

3. 证明 Karamata 定理.

X, X_1, \dots, X_n iid.

(1) $\frac{\sum_{k=1}^n (X_k - m_k)}{a_n} \xrightarrow{d} N(0,1)$.

(2) $L(x) = E[X^2 1_{\{|X| \leq x\}}]$ 在 ∞ 处缓变.

(3) $\lim_{x \rightarrow \infty} \frac{x^2 P(|X| > x)}{E[|X|^2 1_{\{|X| < x\}}]} = 0$

(2) + (3) \Rightarrow (1)?

证: (2) \Rightarrow (3). $\forall r > 1$. (但 r 很接近 1)

选取 $x_0 > 0$ 充分大, 使 $L(2x) \leq r L(x), \forall x \geq x_0$.

对此 $x, x^2 P(|X| > x)$.

$= x^2 \sum_{n=0}^{\infty} P(\frac{|X|}{x} \in (2^n, 2^{n+1}])$

$= x^2 \sum_{n=0}^{\infty} E[1_{\{2^n x < |X| \leq 2^{n+1} x\}}]$

$\stackrel{x^2 \leq |X|^2}{\leq} \sum_{n=0}^{\infty} \frac{1}{4^n} E[|X|^2 1_{\{2^n x < |X| \leq 2^{n+1} x\}}]$

$\leq \sum_{n=0}^{\infty} \frac{1}{4^n} (r-1) r^n L(x) = \frac{1}{1-\frac{r}{4}} L(x)$

$= \frac{(r-1)}{1-\frac{r}{4}} L(x)$

$\therefore \frac{x^2 P(|X| > x)}{L(x)} = \frac{(r-1)}{1-\frac{r}{4}} \rightarrow 0, \text{ as } r \rightarrow 1^+$.

(2) \Rightarrow (1). 证 L 在 ∞ 处缓变.

$L_m(x) := E[(X-m)^2 1_{\{|X-m| \leq x\}}]$

$= E[X^2 1_{\{|X-m| \leq x\}}] + m^2 E[1_{\{|X-m| \leq x\}}]$

$\approx L(x) \therefore L_m(x)$ 在 ∞ 处缓变

证 $E[X] = 0, \forall a_n = 1, \forall \sup\{x > 0; n L(x) \geq x^2\}$
 $a_n \nearrow \infty, n \in \mathbb{N}$

$n L(a_n) \sim a_n^2$

下面先证 (4)

(4) 无

① 无界条件

$$P\left(\left|\frac{X}{a_n}\right| > \varepsilon\right) \sim \frac{a_n^2 P(|X| > a_n \varepsilon)}{n L(a_n)}$$

L 慢变

$$\sim \frac{a_n^2 P(|X| > a_n \varepsilon)}{n \cdot L(a_n \varepsilon)} \rightarrow 0$$

同样

① $n P\left(\left|\frac{X}{a_n}\right| > \varepsilon\right) \rightarrow 0$

② $E X = 0 \therefore n \left[E\left[\frac{X}{a_n} \mathbb{1}_{\left|\frac{X}{a_n}\right| \leq \varepsilon}\right] \right]$

$$\leq \frac{n}{a_n} E\left[|X| \mathbb{1}_{\left|\frac{X}{a_n}\right| > \varepsilon}\right]$$

$$\sim \frac{C_n}{a_n} E\left[|X| \mathbb{1}_{\left|\frac{X}{a_n}\right| > \varepsilon}\right] \rightarrow 0$$

③ $n \text{Var}\left[\frac{X}{a_n} \mathbb{1}_{\left|\frac{X}{a_n}\right| \leq \varepsilon}\right]$

$$= \frac{n}{a_n^2} L(a_n) - n E^2\left[\frac{X}{a_n} \mathbb{1}_{\left|\frac{X}{a_n}\right| \leq \varepsilon}\right] \rightarrow 1$$

\therefore 由 L-条件

□

3. 设 $\{X_n, n \in \mathbb{N}\}$ 独立, 且 $P(X_n = \pm 2^n) = \frac{1}{2^{n+1}}$

$$P(X_n = \pm 1) = \frac{1}{2} \left(1 - \frac{1}{2^n}\right) \quad n \geq 1$$

证明: X_n 不符合 Lindeberg 条件. \int 且从中心极限定理

pf: ① 不合 Lindeberg 条件:

$$\sum_{k=1}^n E\left[X_k^2 \mathbb{1}_{\left|\frac{X_k}{b_k}\right| > \varepsilon}\right] = \sum_{k=1}^n E\left[2^{2k} \cdot \frac{1}{2^k}\right] \rightarrow \infty$$

② 从 CLT: $\forall Y_n = \text{sgn}(X_n) = \begin{cases} 1 \\ -1 \end{cases}$ with prob. $\frac{1}{2}$ each, 只有有限项

$$P(Y_n \neq X_n) = \frac{1}{2^n} \therefore \sum_{n=1}^{\infty} P(Y_n \neq X_n) < \infty \Rightarrow P(Y_n \neq X_n \text{ i.o.}) = 0$$

$E Y_n = 0, E Y_n^2 = \text{Var} Y_n = 1$. ∇ CLT 及 Slutsky 定理

$$\frac{\sum_{k=1}^n X_k}{\sqrt{n}} \xrightarrow{d} N(0,1) \text{ as } n \rightarrow \infty \quad \square$$

4. X_n 独立, $B_n^2 = \sum_{k=1}^n \text{Var}(X_k) \rightarrow \infty$
 $|X_n| \leq C_n = o(B_n)$, 证明: $\frac{S_n - E S_n}{B_n} \xrightarrow{d} N(0,1)$

证明: $\forall Y_k = \frac{X_k - E X_k}{B_n}$

$$\sum_{k=1}^n E Y_k = 0, \sum_{k=1}^n \text{Var} Y_k = \frac{\sum_{k=1}^n \text{Var} X_k}{B_n^2} = 1$$

$\forall \varepsilon > 0$

$$\sum_{k=1}^n E\left[Y_k^2 \mathbb{1}_{\left|\frac{X_k - E X_k}{B_n}\right| > \varepsilon}\right] = \sum_{k=1}^n \left[\frac{1}{B_n^2} E\left[|X_k - E X_k|^2 \mathbb{1}_{\left|\frac{X_k - E X_k}{B_n}\right| > \varepsilon}\right] \right]$$

n 充分大

$$|X_k - E X_k| \leq 2C_k = o(B_n) \text{ as } n \rightarrow \infty$$

$$\therefore \exists N \text{ s.t. } \forall k > N, |X_k - E X_k| < B_n \varepsilon$$

$$\therefore 0 = \sum_{k=1}^n E\left[\frac{1}{B_n^2} (X_k - E X_k)^2\right] \rightarrow 0 \text{ as } n \rightarrow \infty$$

$\therefore Y_n$ 满足 L-条件, 由 CLT 即得

□

5. Y_1, Y_2, \dots, Y_n iid. $\sim E Y = 0, E Y^2 = 1$

$$\{b_n\} \geq 0, b_n = o(B_n) \text{ as } n \rightarrow \infty$$

$$B_n^2 = \sum_{k=1}^n b_k^2, \text{ 证明: } \frac{1}{B_n} \sum_{k=1}^n b_k X_k \xrightarrow{d} N(0,1)$$

证: $Y_k = \frac{b_k X_k}{B_n}, 1 \leq k \leq n, E Y_k = 0$
 $\sum_{k=1}^n E Y_k^2 = \frac{1}{B_n^2} \sum_{k=1}^n b_k^2 X_k^2 = \frac{\sum_{k=1}^n b_k^2}{B_n^2} = 1$

$$\sum_{k=1}^n E\left[Y_k^2 \mathbb{1}_{\left|\frac{Y_k}{b_k}\right| > \varepsilon}\right] = \frac{1}{B_n^2} \sum_{k=1}^n E\left[b_k^2 X_k^2 \mathbb{1}_{\left|\frac{X_k}{b_k}\right| > \varepsilon}\right]$$

对任意 $n, \forall \varepsilon > 0 \exists N$ s.t. $\forall k > N$

$$|X_k| < \frac{B_n}{b_k} \varepsilon$$

$$\text{as } n \rightarrow \infty, \therefore \sum_{k=1}^n E\left[b_k^2 X_k^2 \mathbb{1}_{\left|\frac{X_k}{b_k}\right| > \frac{B_n}{b_k} \varepsilon}\right] \rightarrow 0$$

□

$\{X_k, Y_n\}$ iid. $EX=0, EX^2=1$

Y_n 是独立 i.i.v. $Y_n \in \mathbb{Z}_+$, $Y_n \xrightarrow{a.s.} \infty$

证: $\frac{\sum_{k=1}^{Y_n} X_k}{\sqrt{Y_n}} \xrightarrow{d} N(0,1)$

证:

$$\left| P\left(\frac{\sum_{k=1}^{Y_n} X_k}{\sqrt{Y_n}} \leq x\right) - \Phi(x) \right|$$

$$= \left| \sum_{m=1}^{\infty} P\left(\frac{\sum_{k=1}^{Y_n} X_k}{\sqrt{Y_n}} \leq x \mid Y_n=m\right) P(Y_n=m) - \Phi(x) \right|$$

$$= \left| \sum_{m=1}^{\infty} \left(P\left(\frac{\sum_{k=1}^m X_k}{\sqrt{m}} \leq x \mid Y_n=m\right) - \Phi(x) \right) P(Y_n=m) \right|$$

$\downarrow \forall \varepsilon > 0, \exists N, \forall n > N, \forall k \leq n, \left| P\left(\frac{\sum_{k=1}^m X_k}{\sqrt{m}} \leq x\right) - \Phi(x) \right| < \varepsilon$

$$\leq \sum_{m=1}^N P(Y_n=m) + \sum_{m=N+1}^{\infty} P(Y_n=m)$$

$\underbrace{\hspace{1cm}}_{< \varepsilon} \quad \underbrace{\hspace{1cm}}_{\in P(Y_n \leq N)}$

令 $n \rightarrow \infty$ 有 $\limsup_{n \rightarrow \infty} \left| P\left(\frac{\sum_{k=1}^{Y_n} X_k}{\sqrt{Y_n}} \leq x\right) - \Phi(x) \right| < \varepsilon$

令 $\varepsilon \rightarrow 0^+$ 即得

§3 补充题:

T 为 Ω 上的非负随机变量, $A \subset \Omega$, 令 $T_A = T \mathbb{1}_A + (+\infty) \mathbb{1}_{A^c}$

若 \mathcal{F} 为行时, $A \in \mathcal{F}_\infty$, 则 T_A 为行时 $\iff A \in \mathcal{F}_T$

证: $\Leftarrow \{T_A \leq t\} = \{T \leq t\} \cap A \in \mathcal{F}_t$

$\Rightarrow A \cap \{T < \infty\} = \{T_A < \infty\} \in \mathcal{F}_{T_A}$

$A \cap \{T < \infty\} \cap \{T_A = T\}$

$\in \mathcal{F}_{T_A} \cap \{T_A = T\} = \mathcal{F}_T \cap \{T_A = T\}$

$\in \mathcal{F}_T$

$\Rightarrow A \in \{T < \infty\} \in \mathcal{F}_T$