

一、

1. 基于 $x_s, s < t$ 和 $y_s, s < t$ 对 y_t 的预测误差比基于 $y_s, s < t$ 对 y_t 预测误差小即 $Var(y_t - E[y_t|y_{t-1}, \dots]) \geq Var(y_t - E[y_t|y_{t-1}, \dots, x_{t-1}, \dots])$ $N^2p + \frac{1}{2}p(p+1)$

2. $-0.5 < a < 1$

3. $\frac{1}{1-\phi_0/2-\dots-\phi_p/2^p}$

4. $\begin{cases} 1 + \phi_1^2 + \dots + \phi_{l-1}^2 & l \leq q \\ 1 + \phi_1^2 + \dots + \phi_q^2 & l > q \end{cases}$

5.12 $ARIMA(0, 1, 1) \times (0, 1, 1)_{12}$

6. $GJR-GARCH$ ($TGARCH$) $EGARCH$

7. $EX_t = \mu, EX_t^2 < +\infty, EX_t X_s$ 只与 $(t-s)$ 有关

8. $\frac{\rho_1(1-\rho_2)}{1-\rho_1^2} = \frac{8}{21}$ $\frac{\rho_2-\rho_1^2}{1-\rho_1^2} = \frac{1}{21}$

9. $\cos\omega(t-s)$

二、

1. $p=1$ $d=1$ $q=1$

$Y_t \sim ARIMA(1, 1, 1)$ $\nabla Y_t \sim ARMA(1, 1)$

$\therefore \nabla Y_t = 10 + 0.5\nabla Y_{t-1} + \varepsilon_t - 0.5\varepsilon_{t-1}$

$\therefore E\nabla Y_t = \frac{10}{1-0.5} = 20$

令 $Var(\nabla Y_t) = \gamma_0$ 两边同乘 ε_t $E\nabla Y_t \varepsilon_t = E\varepsilon_t^2 = 1$

$\therefore Var(\nabla Y_t) = Var(10 + 0.5\nabla Y_{t-1} + \varepsilon_t - 0.5\varepsilon_{t-1})$

$\therefore \gamma_0 = 0.25\gamma_0 + 1 + 0.25 - 2 \times 0.25E\nabla Y_{t-1} \varepsilon_{t-1} = 0.25\gamma_0 + 0.75$

$\therefore \gamma_0 = 1$

2. $EX_t \varepsilon_t = E\varepsilon_t^2 = \sigma^2$

$\begin{cases} \gamma_1 = EX_t X_{t-1} = 0.5\gamma_0 - 0.25\sigma^2 \\ \gamma_0 + 0.25\gamma_0 - \gamma_1 = (1 + 0.25^2)\sigma^2 \end{cases}$

解得 $\gamma_0 = \frac{13}{12}\sigma^2$ $\gamma_1 = \frac{7}{24}\sigma^2$ $\therefore \rho_1 = \frac{\gamma_1}{\gamma_0} = \frac{7}{26}$

又 $\therefore \gamma_k = 0.5\gamma_{k-1}, k \geq 2$ $\therefore \rho_k = 0.5\rho_{k-1}, k \geq 2$

$\therefore \rho_k = \frac{7}{26} \times (\frac{1}{2})^{k-1}, k \geq 2$

3. 易知 $\rho_1 = \frac{\theta}{1+\theta^2}$

$\therefore (1+\theta)^2 = 1 + \theta^2 + 2\theta \geq 0$ $(1-\theta)^2 = 1 + \theta^2 - 2\theta \geq 0$

$\therefore 1 + \theta^2 \geq -2\theta$ $1 + \theta^2 \geq 2\theta$

$\therefore |\rho_1| \leq \frac{1}{2}$

三、

1. (1) $|-0.5| < 1, -0.5 \pm (-1) < 1$, 平稳;

$0.4 < 1$, 可逆。

(2) $\psi_0 = 1$

$\psi_1 = \psi_0 = 1$

$\psi_2 = \psi_1 - 0.5\psi_0 = 1 - 0.5 \times 1 = \frac{1}{2}$

$\psi_3 = \psi_2 - 0.5\psi_1 = \frac{1}{2} - 0.5 \times \frac{1}{2} = \frac{1}{4}$

(3)

$$\begin{aligned}
\epsilon_t &= \frac{1 - B + 0.5B^2}{1 + 0.4B} X_t \\
&= (1 - B + 0.5B^2) \sum_{j=0}^{\infty} (-0.4B^j) X_t \\
&= \left(\sum_{j=0}^{\infty} (-0.4)^j B^j - \sum_{j=0}^{\infty} (-0.4)^j B^{j+1} + \sum_{j=0}^{\infty} 0.5(-0.4)^j B^{j+2} \right) X_t \\
&=
\end{aligned}$$

2.(1) 记 $u_t = y_t^2 - h_t$

$$\begin{aligned}
h_t + y_t^2 &= \alpha_0 + \alpha_1 y_{t-1}^2 + \beta_1 h_{t-1} + y_t^2 \\
&= \alpha_0 + \alpha_1 (y_{t-1}^2 - h_{t-1}) + (\alpha_1 + \beta_1) h_{t-1} + y_t^2 \\
\implies y_t^2 &= \alpha_0 + (\alpha_1 + \beta_1) (h_{t-1} + u_{t-1}) + \alpha_1 (y_{t-1}^2 - h_{t-1}) - (\alpha_1 + \beta_1) u_{t-1} + y_t^2 \\
&= \alpha_0 + (\alpha_1 + \beta_1) y_{t-1}^2 + u_t - \beta_1 u_{t-1}
\end{aligned}$$

下证 $u_t \sim WN(0, \sigma^2)$:

$$Eu_t = Ey_t^2 - Eh_t = E(h_t \epsilon_t^2) - Eh_t = Eh_t E\epsilon_t^2 - Eh_t = 0$$

$$\begin{aligned}
\because Cov(u_t, u_s) &= E(u_t u_s) - Eu_t Eu_s \\
&= E(u_t u_s) \\
&= E[(y_t^2 - h_t)(y_s^2 - h_s)] \\
&= E[h_t h_s (\epsilon_t^2 - 1)(\epsilon_s^2 - 1)]
\end{aligned}$$

若 $s < t$, $Cov(u_t, u_s) = E[h_t h_s (\epsilon_t^2 - 1)(\epsilon_s^2 - 1)] = 0$;若 $s = t$, $Cov(u_t, u_s) = E(h_t^2) E[(\epsilon_t^2 - 1)^2] = 2E(h_t^2) = \text{常数}$,故 u_t 为一白噪声。

$$\because E(y_t^4) = 3Eh_t^2$$

$$\begin{aligned}
E(y_t^4) &= [3 + \frac{6\alpha_1^2}{1 - 2\alpha_1^2 - (\alpha_1 + \beta_1)^2}] E(y_t^2)^2 \\
&= [3 + \frac{6\alpha_1^2}{1 - 2\alpha_1^2 - (\alpha_1 + \beta_1)^2}] (\frac{\alpha_0}{1 - \alpha_1 - \beta_1})^2 \\
&\triangleq M
\end{aligned}$$

$$\Rightarrow Eh_t^2 = \frac{M}{3}$$

(2) 由于 $Eh_t = \alpha_0 + \alpha_1 Ey_{t-1}^2 + \beta_1 Eh_{t-1}$

$$\text{且 } Ey_{t-1}^2 = Eh_{t-1}$$

$$\Rightarrow Ey_{t-1}^2 = \frac{\alpha_0}{1 - \alpha_1 - \beta_1}$$

$$\gamma_k = Cov(y_t^2, y_{t+k}^2) = Cov(h_t \epsilon_t^2, h_{t+k} \epsilon_{t+k}^2)$$

$$\text{若 } k = 0, \quad Var(y_t^2) = M - (\frac{\alpha_0}{1 - \alpha_1 - \beta_1})^2 = [2 + \frac{6\alpha_1^2}{1 - 2\alpha_1^2 - (\alpha_1 + \beta_1)^2}] (\frac{\alpha_0}{1 - \alpha_1 - \beta_1})^2$$

若 $k \geq 1$,

$$\begin{aligned}
\gamma_k &= Cov(y_t^2, y_{t+k}^2) \\
&= Cov(u_t, u_{t+k}) + Cov(h_t, y_{t+k}^2) + Cov(y_t^2, h_{t+k}) - Cov(h_t, h_{t+k}) \\
&= Cov(h_t, h_{t+k} \epsilon_{t+k}^2) + Cov(h_t \epsilon_t^2, h_{t+k}) - Cov(h_t, h_{t+k}) \\
&= Cov(h_t \epsilon_t^2, h_{t+k})
\end{aligned}$$

而由于 y_t^2 为 ARMA(1,1)

$$\therefore \gamma_k - (\alpha_1 + \beta_1)\gamma_{k-1} = \begin{cases} \frac{2M}{3}(-\beta_1), & k = 1 \\ 0, & k > 1 \end{cases}$$

故由此可递推得 $\{\gamma_k\}_{k=0}^\infty$.

3.(1) 由 ARMA 定义知 $E(X_s, \epsilon_t) = 0, \forall s < t$, 且 $\epsilon_t \sim WN(O, \sigma^2)$.

故 $Cov(X_t, \epsilon_t) = Cov(\phi_1 X_{t-1} + \epsilon_t + \theta_1 \epsilon_{t-1}, \epsilon_t) = \sigma^2$.

$$\begin{aligned}
\gamma_0 &= Cov(X_t, X_t) \\
&= Cov(\phi_1 X_{t-1} + \epsilon_t + \theta_1 \epsilon_{t-1}, \phi_1 X_{t-1} + \epsilon_t + \theta_1 \epsilon_{t-1}) \\
&= \phi_1^2 \gamma_0 + (2\phi_1 \theta_1 + \theta_1^2 + 1)\sigma^2
\end{aligned}$$

$$\Rightarrow \gamma_0 = \frac{2\phi_1 \theta_1 + \theta_1^2 + 1}{1 - \phi_1^2} \sigma^2$$

$$\gamma_1 = Cov(X_t, X_{t-1}) = \gamma_0 + \theta_1 \sigma^2$$

$$\gamma_2 = Cov(X_t, X_{t-2}) = \phi_1 \gamma_1$$

$$\text{且 } \gamma_n = Cov(X_t, X_{t-n}) = \phi_1 \gamma_{n-1} = \phi_1^{n-1} \gamma_1 \quad (\forall n \geq 2)$$

$$\text{故 } \Gamma = \begin{pmatrix} Cov(X_t, X_t) & Cov(X_t, X_{t-1}) \\ Cov(X_t, X_{t-1}) & Cov(X_{t-1}, X_{t-1}) \end{pmatrix} = \begin{pmatrix} \gamma_0 & \gamma_1 \\ \gamma_1 & \gamma_0 \end{pmatrix}$$

$$\beta = \begin{pmatrix} Cov(X_t, Y_{t+1}) \\ Cov(X_{t-1}, Y_{t+1}) \end{pmatrix} = \begin{pmatrix} \frac{1}{2}(\gamma_0 + \gamma_1) \\ \frac{1}{2}(\gamma_1 + \gamma_2) \end{pmatrix}$$

$$\Rightarrow a = \Gamma^{-1} \beta$$

故均方误差为 $a(X_t, X_{t-1})^T$.

(2) 记 $\{Y_t\}$ 的协方差函数为 $\{\hat{\gamma}_k\}$.

$$\text{则 } \{\hat{\gamma}_k\} = Cov(Y_t, Y_{t+k}) = \frac{1}{4}(2\gamma_k + \gamma_{k+1} + \gamma_{k-1}).$$

易知, γ_k 为绝对可和的。

而

$$\begin{aligned}
\sum_{k=-\infty}^{\infty} |\hat{\gamma}_k| &\leq \frac{1}{4} \sum_{k=-\infty}^{\infty} (2|\gamma_k| + |\gamma_{k+1}| + |\gamma_{k-1}|) \\
&= \sum_{k=-\infty}^{\infty} |\gamma_k| \\
&< \infty
\end{aligned}$$

故 $\{Y_t\}$ 有谱密度:

$$f(\lambda) = \frac{1}{2\pi} \sum_{k=-\infty}^{\infty} \hat{\gamma}_k e^{-ik\lambda},$$

$\{\hat{\gamma}_k\}$ 可由(1)中 γ_k 算得。